NPFL129, Lecture 9



Correlation, Model Combination, Decision Trees, Random Forests

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unless otherwise stated

Covariance



Given a collection of random variables x_1,\ldots,x_N , we know that

$$\mathbb{E}\left[\sum_i \mathrm{x}_i
ight] = \sum_i \mathbb{E}ig[\mathrm{x}_iig].$$

But how about $\operatorname{Var}\left(\sum_{i} \mathbf{x}_{i}\right)$?

$$\begin{aligned} \operatorname{Var}\left(\sum_{i} \mathbf{x}_{i}\right) &= \mathbb{E}\left[\left(\sum_{i} \mathbf{x}_{i} - \sum_{i} \mathbb{E}[\mathbf{x}_{i}]\right)^{2}\right] \\ &= \mathbb{E}\left[\left(\sum_{i} \left(\mathbf{x}_{i} - \mathbb{E}[\mathbf{x}_{i}]\right)\right)^{2}\right] \\ &= \mathbb{E}\left[\sum_{i} \sum_{j} \left(\mathbf{x}_{i} - \mathbb{E}[\mathbf{x}_{i}]\right) \left(\mathbf{x}_{j} - \mathbb{E}[\mathbf{x}_{j}]\right)\right] \\ &= \sum_{i} \sum_{j} \mathbb{E}\left[\left(\mathbf{x}_{i} - \mathbb{E}[\mathbf{x}_{i}]\right) \left(\mathbf{x}_{j} - \mathbb{E}[\mathbf{x}_{j}]\right)\right].\end{aligned}$$

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Covariance



We define **covariance** of two random variables x, y as

$$Cov(\mathbf{x}, \mathbf{y}) = \mathbb{E}\left[\left(\mathbf{x} - \mathbb{E}[\mathbf{x}]\right)\left(\mathbf{y} - \mathbb{E}[\mathbf{y}]\right)\right].$$
$$V_{\mathcal{W}}(\mathbf{x}) = \mathbb{E}(\mathbf{x}^{2}) - \mathbb{E}(\mathbf{x})^{2}$$
$$Var\left(\sum_{i} \mathbf{x}_{i}\right) = \sum_{i} \sum_{j} Cov(\mathbf{x}_{i}, \mathbf{x}_{j}).$$

Then,

Note that Cov(x,x) = Var(x) and that we can write covariance as

$$\begin{split} \mathrm{Cov}(\mathbf{x},\mathbf{y}) &= \mathbb{E}\Big[\big(\mathbf{x} - \mathbb{E}[\mathbf{x}]\big)\big(\mathbf{y} - \mathbb{E}[\mathbf{y}]\big)\Big] \\ &= \mathbb{E}\big[\mathbf{x}\mathbf{y} - \mathbf{x}\mathbb{E}[\mathbf{y}] - \mathbb{E}[\mathbf{x}]\mathbf{y} + \mathbb{E}[\mathbf{x}]\mathbb{E}[\mathbf{y}]\big] \\ &= \mathbb{E}\big[\mathbf{x}\mathbf{y}\big] - \mathbb{E}\big[\mathbf{x}\big]\mathbb{E}\big[\mathbf{y}\big]. \end{split}$$

Correlation



Two random variables x, y are **uncorrelated** if Cov(x, y) = 0; otherwise, they are **correlated**.

Note that two *independent* random variables are uncorrelated, because

$$\begin{aligned} \operatorname{Cov}(\mathbf{x},\mathbf{y}) &= \mathbb{E}\left[\left(\mathbf{x} - \mathbb{E}[\mathbf{x}]\right)\left(\mathbf{y} - \mathbb{E}[\mathbf{y}]\right)\right] \\ &= \sum_{x,y} \underbrace{P(x,y)\left(x - \mathbb{E}[x]\right)\left(y - \mathbb{E}[y]\right)}_{x,y} \underbrace{P(x)\left(x - \mathbb{E}[x]\right)P(y)\left(y - \mathbb{E}[y]\right)}_{y} \\ &= \sum_{x,y} P(x)\left(x - \mathbb{E}[x]\right)\sum_{y} P(y)\left(y - \mathbb{E}[y]\right) \\ &= \sum_{x} P(x)\left(x - \mathbb{E}[x]\right)\sum_{y} P(y)\left(y - \mathbb{E}[y]\right) \\ &= \mathbb{E}_{\mathbf{x}}\left[\mathbf{x} - \mathbb{E}[\mathbf{x}]\right]\mathbb{E}_{\mathbf{y}}\left[\mathbf{y} - \mathbb{E}[y]\right] = 0. \end{aligned}$$

However, dependent random variables can be uncorrelated – random uniform x on [-1, 1] and y = |x| are not independent (y is completely determined by x), but they are uncorrelated.

There are several ways to measure correlation of random variables x, y.

Pearson correlation coefficient, denoted as ρ or r, is defined as

$$ho \stackrel{ ext{def}}{=} rac{ ext{Cov}(\mathrm{x},\mathrm{y})}{\sqrt{ ext{Var}(\mathrm{x})}\sqrt{ ext{Var}(\mathrm{y})}} \int dh_{2} ext{truth} ext{truth} ext{truth} ext{truth} ext{eleft} = rac{1}{\sqrt{ ext{Var}(\mathrm{x})}\sqrt{ ext{Var}(\mathrm{y})}}{r \stackrel{ ext{def}}{=} rac{\sum_{i}(x_{i}-ar{x})(y_{i}-ar{y})}{\sqrt{\sum_{i}(x_{i}-ar{x})^{2}}\sqrt{\sum_{i}(y_{i}-ar{y})^{2}}},$$

where:

- ρ is used when the full expectation is computed (population Pearson correlation coefficient);
- r is used when estimating the coefficient from data (sample Pearson correlation coefficient);
 - $\circ~ar{x}$ and $ar{y}$ are sample estimates of the respective means.

The value of Pearson correlation coefficient is in fact normalized covariance, because its value is always bounded by $-1 \le \rho \le 1$ (and the same holds for r).

The bound can be derived from

$$egin{aligned} &0 \leq \mathbb{E}iggl[iggl(rac{(\mathrm{x}-\mathbb{E}[\mathrm{x}])}{\sqrt{\mathrm{Var}(\mathrm{x})}} -
horac{(\mathrm{y}-\mathbb{E}[\mathrm{y}])}{\sqrt{\mathrm{Var}(\mathrm{y})}}iggr)^2iggr] \ &= \mathbb{E}iggl[rac{(\mathrm{x}-\mathbb{E}[\mathrm{x}])^2}{\mathrm{Var}(\mathrm{x})}iggr] - 2
ho\mathbb{E}iggl[rac{(\mathrm{x}-\mathbb{E}[\mathrm{x}])}{\sqrt{\mathrm{Var}(\mathrm{x})}}rac{(\mathrm{y}-\mathbb{E}[\mathrm{y}])}{\sqrt{\mathrm{Var}(\mathrm{y})}}iggr] +
ho^2\mathbb{E}iggl[rac{(\mathrm{y}-\mathbb{E}[\mathrm{y}])^2}{\mathrm{Var}(\mathrm{y})}iggr] \ &= rac{\mathrm{Var}(\mathrm{x})}{\mathrm{Var}(\mathrm{x})} - 2
ho \cdot
ho +
ho^2rac{\mathrm{Var}(\mathrm{y})}{\mathrm{Var}(\mathrm{y})} = 1 -
ho^2, \end{aligned}$$

which yields $ho^2 \leq 1.$

Alternatively, the desired inequality can be obtained by applying the Cauchy-Schwarz inequality $\langle u,v
angle \leq \sqrt{\langle u,u
angle}\sqrt{\langle v,v
angle}$ on $\langle x,y
angle \stackrel{\text{\tiny def}}{=} \mathbb{E}[xy]$.



Pearson correlation coefficient quantifies linear dependence of the two random variables.



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Pearson correlation coefficient quantifies linear dependence of the two random variables.



The four displayed variables have the same mean 7.5, variance 4.12, Pearson correlation coefficient 0.816 and regression line $3 + \frac{1}{2}x$.



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Nonlinear Correlation – Spearman's ρ

To measure also nonlinear correlation, two coefficients are commonly used.

Spearman's rank correlation coefficient ρ

Spearman's ρ is Pearson correlation coefficient measured on **ranks** of the original data, where a rank of an element is its index in sorted ascending order.



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Nonlinear Correlation – Kendall's τ

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Kendall rank correlation coefficient τ

Kendall's τ measures the amount of *concordant pairs* (pairs where y increases/decreases when x does), minus the *discordant pairs* (where y increases/decreases when x does the opposite):

There is no clear consensus whether to use Spearman's ρ or Kendall's τ . When there are no/few ties in the data, Kendall's τ offers two minor advantages $-\frac{1+\tau}{2}$ can be interpreted as a probability of a concordant pair, and Kendall's τ converges to a normal distribution faster. As defined, the range of Kendall's $\tau \in [-1, 1]$. However, if there are ties, its range is smaller – therefore, several corrections (not discussed here) exist to adjust its value in case of ties.

Usage of Correlation in Machine Learning

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In the machine learning area, correlation is commonly used to measure "agreement" between:

- several human annotations;
- automatic metric and gold annotation;



• automatic metric and human evaluation.



Model Combination aka Ensembling

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The goal of **ensembling** is to combine several models in order to reach higher performance.

The simplest approach is to train several independent models and then to combine their predictions by averaging or voting.

The terminology varies, but for classification:

- voting (or hard voting) usually means predicting the class predicted most often by the individual models,
- averaging (or soft voting) denotes averaging the returned model distributions and predicting the class with the highest probability.

The main idea behind ensembling is that if models have uncorrelated errors, then by averaging model predictions the errors will cancel out.

Visualization of Ensembling Performance

Consider ensembling predictions generated uniformly on a planar disc:



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Model Combination aka Ensembling

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If we denote the prediction of a model y_i on a training example (\boldsymbol{x}, t) as $y_i(\boldsymbol{x}) = t + \varepsilon_i(\boldsymbol{x})$, so that $\varepsilon_i(\boldsymbol{x})$ is the model error on example \boldsymbol{x} , the mean square error of the model is

$$\mathbb{E}ig[(y_i(oldsymbol{x})-t)^2ig] = \mathbb{E}ig[arepsilon_i^2(oldsymbol{x})ig].$$

Considering M models, we analogously get that the mean square error of the ensemble is

$$\mathbb{E}igg[\Big(rac{1}{M} \sum_i arepsilon_i(oldsymbol{x}) \Big)^2 igg].$$

Finally, assuming that the individual errors ε_i have zero mean and are *uncorrelated*, we get that $\mathbb{E}[\varepsilon_i(\boldsymbol{x})\varepsilon_j(\boldsymbol{x})] = 0$ for $i \neq j$, and therefore, $\mathbb{E}[\left(\frac{1}{M}\sum_i \varepsilon_i(\boldsymbol{x})\right)^2] = \mathbb{E}\left[\frac{1}{M^2}\sum_{i,j}\varepsilon_i(\boldsymbol{x})\varepsilon_j(\boldsymbol{x})\right] = \frac{1}{M}\mathbb{E}\left[\frac{1}{M}\sum_i \varepsilon_i^2(\boldsymbol{x})\right],$ so the average error of the ensemble is $\frac{1}{M}$ times the average error of the individual models. **NPFL129, Lecture 9** Covariance Correlation Model Combination Decision Tree Gin and Entropy Losses RandomForests 15/25

Bagging – Bootstrap Aggregation

For neural network models, training models with independent random initialization is usually enough, given that the loss has many local minima, so the models tend to be quite independent just when using different random initialization.

However, algorithms with convex loss functions usually converge to the same optimum independent of randomization.

In these cases, we can use **bagging**, which stands for **bootstrap aggregation**.

In bagging, we construct a different dataset for every model to be trained. We construct it using **bootstrapping** – we sample as many training instances as the original dataset has, but **with replacement**.

Such dataset is sampled using the same empirical data distribution and has the same size, but is not identical.





Figure 7.5 of "Deep Learning" book, https://www.deeplearningbook.org

DecisionTree

Decision Trees

The idea of decision trees is to partition the input space into usually cuboid regions and solving each region with a simpler model.

We focus on **Classification and Regression Trees** (CART; Breiman et al., 1984), but there are additional variants like ID3, C4.5, ... bod product in both product in the second product product in the second product in the second product produc



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Regression Decision Trees

Assume we have an input dataset $X \in \mathbb{R}^{N \times D}$, $t \in \mathbb{R}^N$. At the beginning, the decision tree is just a single node and all input examples belong to this node. We denote $I_{\mathcal{T}}$ the set of training example indices belonging to a node \mathcal{T} .

For each leaf (a node without children), our model predicts the average of the training examples belonging to that leaf, $\hat{t}_{\mathcal{T}} = \frac{1}{|I_{\mathcal{T}}|} \sum_{i \in I_{\mathcal{T}}} t_i$. If the training examples of the training examples of a node \mathcal{T} are - for regression, we employ the sum of squares error between the examples belonging to the node and the predicted value in that node; this is proportional to the variance of the training examples. Note that even if it is not mean squared error, it is sometimes denoted as MSE.

$$c_{ ext{SE}}(\mathcal{T}) \stackrel{ ext{\tiny def}}{=} \sum_{i \in I_\mathcal{T}} (t_i - \hat{t}_\mathcal{T})^2, \,\, ext{where} \,\, \hat{t}_\mathcal{T} = rac{1}{|I_\mathcal{T}|} \sum_{i \in I_\mathcal{T}} t_i \,,$$

Tree Construction

To split a node, the goal is to find a feature and its value such that when splitting a node \mathcal{T} into \mathcal{T}_L and \mathcal{T}_R , the resulting regions decrease the overall criterion value the most, i.e., the difference $c_{\mathcal{T}_L} + c_{\mathcal{T}_R} - c_{\mathcal{T}}$ is the lowest. - chei to note the help help V listed co chei what when splitting constraints, the most common ones are: negadobhajsi data.

- maximum tree depth: we do not split nodes with this depth;
- minimum examples to split: we only split nodes with this many training examples;
- maximum number of leaf nodes: we split until we reach the given number of leaves.

The tree is usually built in one of two ways:

- if the number of leaf nodes is unlimited, we usually build the tree in a depth-first manner, recursively splitting every leaf until one of the above constraints is invalidated;
- if the maximum number of leaf nodes is given, we usually split such leaf \mathcal{T} where the criterion difference $c_{\mathcal{T}_L} + c_{\mathcal{T}_R} c_{\mathcal{T}}$ is the lowest.

Hrvení jednoduše dverfitným, že udělám příliš větní.

Covariance

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RandomForests

To control overfitting, the previously-mentioned constraints can be used.

Additionally, **pruning** can also be used. After training, we might decide that some subtrees are not necessary, and *prune* them (replacing them with a leaf). Pruning can be used both as a regularization or model compression.

There are many heuristics to prune a decision tree; Scikit-learn implements **minimal costcomplexity pruning**:

• we extend the criterion to *cost-complexity criterion* as

$$r_{lpha}$$
 for a leaf, $c_{lpha}(\mathcal{T})=c(\mathcal{T})+lpha$,

 $^{\circ}$ for a subtree with a root t, $c_{lpha}(t) = \sum_{\mathcal{T} \in ext{leaves}} c_{lpha}(\mathcal{T}) = \sum_{\mathcal{T} \in ext{leaves}} c(\mathcal{T}) + lpha | ext{leaves}|;$

- generally a criterion in a node t is greater or equal to the sum of criteria of its leaves;
- α_{eff} for a node t is the value of α such that the two mentioned cost-complexity quantities ($c_{\alpha}(\mathcal{T})$ computed as if t is a leaf, and $c_{\alpha}(t)$) are equal $\circ \alpha_{\text{eff}} = (c(\mathcal{T}) - c(t))/(|\text{leaves under }t| - 1);$
- we then prune the nodes in the order of increasing $lpha_{
 m eff}.$



Classification Decision Trees

For multi-class classification, we predict the class which is the most frequent in the training examples belonging to a leaf \mathcal{T} .

To define the criteria, let us denote the average probability for class k in a region \mathcal{T} as $p_{\mathcal{T}}(k)$.

For classification trees, one of the following two criteria is usually used:

• Gini index, also called Gini impurity, measuring how often a randomly chosen element would be incorrectly labeled if it was randomly labeled according to p_T :

$$c_{ ext{Gini}}(\mathcal{T}) \stackrel{ ext{\tiny def}}{=} |I_\mathcal{T}| \sum_k p_\mathcal{T}(k) ig(1-p_\mathcal{T}(k)ig),$$

• Entropy Criterion

$$c_{ ext{entropy}}(\mathcal{T}) \stackrel{ ext{def}}{=} |I_\mathcal{T}| \cdot H(oldsymbol{p}_\mathcal{T}) = - |I_\mathcal{T}| \sum_{\substack{k \ p_\mathcal{T}(k)
eq 0}} p_\mathcal{T}(k) \log p_\mathcal{T}(k).$$

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DecisionTree



Binary Gini as (M)SE Loss

Recall that $I_{\mathcal{T}}$ denotes the set of training example indices belonging to a leaf node \mathcal{T} , let $n_{\mathcal{T}}(0)$ be the number of examples with target value 0, $n_{\mathcal{T}}(1)$ be the number of examples with target value 1, and let $p_{\mathcal{T}} = \frac{1}{|I_{\mathcal{T}}|} \sum_{i \in I_{\mathcal{T}}} t_i = \frac{n_{\mathcal{T}}(1)}{n_{\mathcal{T}}(0) + n_{\mathcal{T}}(1)}$.

Consider sum of squares loss $L(p) = \sum_{i \in I_T} (p - t_i)^2$.

By setting the derivative of the loss to zero, we get that the p minimizing the loss fulfills $|I_T|p = \sum_{i \in I_T} t_i$, i.e., $p = p_T$.

The value of the loss is then

Covariance

$$egin{split} L(p_{\mathcal{T}}) &= \sum_{i \in I_{\mathcal{T}}} (p_{\mathcal{T}} - t_i)^2 = n_{\mathcal{T}}(0)(p_{\mathcal{T}} - 0)^2 + n_{\mathcal{T}}(1)(p_{\mathcal{T}} - 1)^2 \ &= rac{n_{\mathcal{T}}(0)n_{\mathcal{T}}(1)^2}{\left(n_{\mathcal{T}}(0) + n_{\mathcal{T}}(1)
ight)^2} + rac{n_{\mathcal{T}}(1)n_{\mathcal{T}}(0)^2}{\left(n_{\mathcal{T}}(0) + n_{\mathcal{T}}(1)
ight)^2} = rac{(n_{\mathcal{T}}(1) + n_{\mathcal{T}}(0))n_{\mathcal{T}}(0)n_{\mathcal{T}}(1)}{\left(n_{\mathcal{T}}(0) + n_{\mathcal{T}}(1)
ight)\left(n_{\mathcal{T}}(0) + n_{\mathcal{T}}(1)
ight)} \ &= ig(n_{\mathcal{T}}(0) + n_{\mathcal{T}}(1)ig)(1 - p_{\mathcal{T}})p_{\mathcal{T}} = |I_{\mathcal{T}}| \cdot p_{\mathcal{T}}(1 - p_{\mathcal{T}}). \end{split}$$

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Gini and Entropy Losses

Entropy as NLL Loss



Again let $I_{\mathcal{T}}$ denote the set of training example indices belonging to a leaf node \mathcal{T} , let $n_{\mathcal{T}}(k)$ be the number of examples with target value k, and let $p_{\mathcal{T}}(k) = \frac{1}{|I_{\mathcal{T}}|} \sum_{i \in I_{\mathcal{T}}} [t_i = k] = \frac{n_{\mathcal{T}}(k)}{|I_{\mathcal{T}}|}$. Consider a distribution $m{p}$ on K classes and non-averaged NLL loss $L(m{p}) = \sum_{i \in I_{\tau}} -\log p_{t_i}$. By setting the derivative of the loss with respect to p_k to zero (using a Lagrangian with constraint $\sum_k p_k = 1$), we get that the $m{p}$ minimizing the loss fulfills $p_k = p_{\mathcal{T}}(k)$. $\mathcal{N}_{\mathsf{T}}(\mathsf{L}) = \rho_{\mathsf{T}}(\mathsf{L}) \cdot |\mathcal{I}_{\mathsf{T}}|$ The value of the loss with respect to $p_{\mathcal{T}}$ is then $L(oldsymbol{p}_{\mathcal{T}}) = \sum -\log p_{t_i}$ $i \in I_{\mathcal{T}}$ $= - \sum n_{\mathcal{T}}(k) \log p_{\mathcal{T}}(k)$ $\substack{k \ p_{\mathcal{T}}(k)
eq 0}$ $|z| = -|I_\mathcal{T}| \quad \sum \quad p_\mathcal{T}(k) \log p_\mathcal{T}(k) = |I_\mathcal{T}| \cdot H(oldsymbol{p}_\mathcal{T}).$ $k \\ p_{\mathcal{T}}(k) \neq 0$

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Random Forests



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Bagging of data combined with a random subset of features (sometimes called *feature bagging*).



Random Forests



Bagging

Every decision tree is trained using bagging (on a bootstrapped dataset).

Random Subset of Features

During each node split, only a random subset of features is considered when finding the best split. A fresh random subset is used for every node.

Extra Trees

The so-called extra trees are even more randomized, not finding the best possible feature value when choosing a split, but considering uniformly random samples from a feature's empirical range (minimum and maximum in the training data).

Demo

https://cs.stanford.edu/~karpathy/svmjs/demo/demoforest.html